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ABSTRAK

Perceraian merupakan fenomena sosial yang terus meningkat dan menjadi perhatian dalam dinamika kehidupan keluarga di Indonesia, termasuk di Provinsi Kalimantan Barat. Penelitian ini bertujuan untuk menganalisis pengaruh perkawinan usia dini, rata-rata lama sekolah, dan tingkat kemiskinan terhadap jumlah perceraian di Kalimantan Barat. Metode yang digunakan adalah pendekatan kuantitatif dengan analisis regresi data panel menggunakan data sekunder periode 2018–2024 yang bersumber dari Badan Pusat Statistik dan laporan Pengadilan Agama se-Kalimantan Barat. Model estimasi terbaik yang diperoleh melalui uji Chow dan uji Hausman adalah model fixed effect. Hasil penelitian menunjukkan bahwa tingkat kemiskinan berpengaruh positif dan signifikan terhadap jumlah perceraian, hal ini berarti peningkatan kemiskinan cenderung meningkatkan jumlah perceraian. Perkawinan usia dini juga berpengaruh positif terhadap perceraian. Sementara itu, rata-rata lama sekolah berpengaruh negatif dan signifikan terhadap jumlah perceraian, yang menunjukkan bahwa peningkatan tingkat pendidikan dapat menurunkan risiko perceraian. Nilai R-squared sebesar 0,921393 merupakan ukuran goodness-of-fit model fixed effect dan tidak diatribusikan seluruhnya kepada tiga variabel penjelas karena mencakup kontribusi efek tetap. Temuan ini menunjukkan bahwa faktor sosial ekonomi memiliki peran penting dalam stabilitas rumah tangga, sehingga peningkatan kualitas pendidikan dan perbaikan kondisi ekonomi masyarakat menjadi langkah strategis dalam menekan angka perceraian.

Kata kunci: perceraian, kemiskinan, rata-rata lama sekolah, usia pernikahan pertama, regresi data panel

ABSTRACT

Divorce has increasingly become a significant social issue affecting family dynamics in Indonesia, including in West Kalimantan Province. This study aims to examine the influence of early marriage, mean years of schooling, and poverty rate on the number of divorces in West Kalimantan. The study applies a quantitative approach using panel data regression based on secondary data from 2018–2024 obtained from Statistics Indonesia and reports from Religious Courts across West Kalimantan. The most appropriate estimation model, determined through Chow and Hausman tests, is the fixed-effects model. The findings indicate that poverty has a positive and statistically significant effect on divorce, suggesting that higher levels of poverty are associated with a greater likelihood of marital dissolution. Early marriage also shows a positive relationship with the number of divorces. In contrast, the mean years of schooling demonstrate a negative and significant effect on divorce, implying that higher educational attainment tends to reduce the probability of divorce. The R-squared value of 0.921393 serves as the

goodness-of-fit measure for the fixed effect model and is not entirely attributed to the three explanatory variables, as it incorporates the contribution of fixed effects.

These results highlight the crucial role of socio-economic conditions in maintaining household stability. Strengthening educational attainment and improving economic well-being may therefore serve as effective strategies to reduce number of divorces.

Keywords: divorce, poverty, mean years of schooling, early marriage, panel data regression

INTRODUCTION

Divorce has emerged as an increasingly prevalent social issue across various regions in Indonesia. In West Kalimantan, this phenomenon warrants critical attention due to its profound ramifications on the well-being of both individuals and family structures. Stevenson and Wolfers (2007) show that changes in economic structure and marriage market characteristics are an important part of explaining shifts in marriage and divorce patterns. While divorce can occasionally yield positive outcomes such as liberating individuals from severe domestic conflict or prolonged emotional distress it predominantly inflicts adverse effects on the individuals involved, further expanding into broader familial, social, and economic disruptions (Nadiatusholikha et al., 2024). Fundamentally, every individual aspires to cultivate a harmonious family, given that the household serves as the primary environment for early human socialization. The role of a stable family unit is pivotal in shaping a child's future,

acting as the baseline nurturing environment that guarantees their fundamental rights to sanctuary and lifelong security. Conversely, when a family fails to execute this nurturing function, children are highly susceptible to various maladaptive outcomes. These include drifting into street life, joining marginalized or homeless subcultures where they face severe neglect, and developing psychosocial disorders that ultimately impair their academic performance (Mone, 2019).

Corroborating this, research by Sindarti et al. (2019) demonstrated that children raised in broken homes frequently suffer from diminished learning motivation, higher truancy rates, and unmonitored behavior outside the school environment. This domestic instability also triggers school dropout rates, as children from broken homes often lose the enthusiasm to pursue education or face a complete deficit in financial support to sustain their schooling. Essentially, a family is designed to foster a sense of belonging, safety, affection, and mutual harmony among its members. The emotional bond

within a family extends beyond mere sentiments; it encompasses sustained care, mutual responsibility, attentiveness, deep understanding, respect, and a shared commitment to fostering optimal child growth and development (Mone, 2019).

According to empirical data compiled from the Statistics Indonesia (BPS) website for West Kalimantan Province spanning 2018 to 2024, there is a discernable downward trend in marriages alongside a concurrent rise in number of divorces across nearly all religious courts in the region. In 2018, West Kalimantan recorded 25,660 marriages against 7,060 divorce cases. Consequently, the divorce-to-marriage ratio for that year stood at approximately 27 percent. The highest proportion of divorces over this seven-year period occurred in 2022, spiking to roughly 41 percent of the total marriages registered that year. Furthermore, a notable decline in marriage registration occurred in 2024, dropping to 21,064 incidents from 21,533 in 2023. This contraction might stem from a growing "marriage phobia" within society, driven by the

apprehension that marital unions inevitably end in dissolution, thereby discouraging individuals from marrying altogether. This trend poses a severe macroeconomic threat; as marriage rates dwindle, the sustainability of future generations is compromised, precisely at a time when Indonesia heavily anticipates capitalizing on its demographic dividend. These changing household dynamics are likely multi-causal, influenced by shifts in social values, heightened financial distress, lack of marital readiness among young couples, and the lingering socioeconomic shocks of the COVID-19 pandemic since early 2020.

The early marriage significantly dictates marital longevity, as cognitive maturity and problem-solving capacities are vital for maintaining relationship stability. Ideally, a husband must be prepared to assume responsibility as the household head, while a wife must be ready to manage the domestic sphere, nurture, and educate the offspring (Afifatimah et al., 2023). To address this, the government has instituted legal age thresholds for marriage through various

considerations, as codified in Law No. 16 of 2019 Article 7 Paragraph (1), which mandates that marriage is only permissible once both men and women reach the age of 19 (Heriyani et al., 2022). Widyastari et al. (2020) found that the relationship between age at marriage and marital dissolution is an important issue in Indonesia. Meanwhile, Garcia-Hombrados and Özcan (2024) found that marrying at a younger age (under 20 or in their early 20s) carries a higher risk of divorce. Research by Cameron et al. (2023) found that men who marry at an early age have a higher probability of experiencing divorce by about 5 percentage points, or an increase of more than 100% compared to men who marry at an older age. Nevertheless, early marriage is frequently driven by parental economic constraints; parents often perceive that marrying off a child absolves them of educational and financial burdens. This aligns with empirical findings by Badruzaman (2021), Hidayat et al. (2025) and Haque & Irmawati (2025), which indicate that through marriage, the socioeconomic responsibility over the child officially

shifts to the husband. Hence, a clear correlation exists between parental socioeconomic vulnerability and the prevalence of youth marriage.

Data on early marriages in West Kalimantan reveals volatile dynamics over the last seven years. The number of cases escalated drastically in 2020, jumping from 506 cases in 2019 to 1,253 cases, and remained elevated in 2021 at 1,216 cases. This surge can be tied to the pandemic-induced social and economic disruptions, which altered adolescent activity patterns, amplified familial financial pressures, and restricted educational access. Conversely, a contraction was observed between 2022 and 2024, where cases dropped from 959 to 547. This decline reflects positive progress, presumably driven by growing public awareness, stricter enforcement of minimum age regulations, and the efficacy of child marriage prevention initiatives.

Beyond age maturity, educational attainment significantly alters how an individual perceives conflict and exercises discretion in decision-making. Education also serves as a primary

determinant of employment opportunities. Generally, higher educational attainment unlocks superior professional positions, which ultimately enhances an individual's economic standing (Dyah et al., 2023). Economically secure households foster comfort and marital satisfaction, thereby driving down number of divorces. In contrast, unemployment of the household head generates severe economic strain, triggering chronic domestic disputes that erode emotional bonds and heighten the risk of marital dissolution (Rabbani et al., 2025).

The mean years of schooling (MYS) in West Kalimantan from 2018 to 2024 demonstrated a steady, albeit decelerating, increase. This upward trajectory indicates a gradual improvement in public education quality. In the context of marital life, higher education can mitigate divorce risks because educated couples typically exhibit superior communication skills, better conflict resolution mechanisms, and more structured family planning. However, if educational gains are not supported by economic stability and

emotional quality within the household, divorce risks remain pertinent.

Alongside marital age and education, broader economic factors heavily dictate domestic stability (Abubakar, 2023). Among these, poverty stands out as a core catalyst for divorce. Although the impoverished population in West Kalimantan has shown a declining trend over the past five years, the absolute number of poor individuals remains substantial, spanning tens to hundreds of thousands.

Consequently, macroeconomic shocks such as inflation, food crises, or mass layoffs can easily push these vulnerable groups back below the poverty line. Impoverished communities are structurally characterized by financial incapacity. Extensive literature (Azis, 2021; Manna, 2021; Nadiatusholikha et al., 2024; Rabbani et al., 2025) underscores that regional poverty rates are intrinsically tied to family economic instability, which frequently serves as the bedrock for divorce.

RESEARCH METHODOLOGY

This study employs a quantitative method, focusing on West Kalimantan Province as the research locus. The analysis utilizes secondary data

spanning the 2018–2024 period, sourced from Statistics Indonesia (BPS) of West Kalimantan Province and official reports from Religious Courts across regencies and cities throughout West Kalimantan.

Variabel	Keterangan	Deskripsi
PCR	Number of divorces (Unit: Case)	The number of divorces cases across regencies/cities by year
EM	Early marriage (Unit: Case)	The number of early marriage cases recorded across regions by year
MYS	Mean Years of Schooling (Unit: Year)	The average number of years of formal education completed by the population aged 15 years and above in a province, based on annual BPS Susenas data
Pov	Percentage of Poor Population (Unit: Percent)	The percentage of the population living below the official poverty line set by BPS, measured based on monthly per capita expenditure consumption. Data source is obtained from annual BPS publications at the provincial level

Source: BPS

The empirical model used in this study is formulated as follows:

$$PCR = \beta_0 + \beta_1 EM_{it} + \beta_2 MYS_{it} + \beta_3 POV_{it} + \varepsilon_{it}$$

Where:

PCR = Number of divorces

EM = Early marriage cases

RLS = Mean years of schooling

POV = Poverty rate

This study utilizes a panel data regression structure equation as the

primary analysis tool. Panel data is a combination of cross-section data and time-series data. Cross-section data refers to data on one or more variables collected for multiple individuals at a single point in time. Meanwhile, time-series data consists of data on one or more variables collected sequentially over time. Therefore, in panel data, the same cross-sectional units are observed and collected over a specific period of time (Gujarati, 2004).

In this study, classical assumption tests and Robustness Check were conducted. A classical assumption test is a series of statistical tests performed on a regression model (typically multiple linear regression) to ensure that the model satisfies the fundamental requirements of the Ordinary Least Squares (OLS) method, thereby ensuring that the estimated results are valid, unbiased, and reliable. The classical assumption tests performed in this research consist of the normality test, multicollinearity test, autocorrelation test, and heteroscedasticity test. Model robustness testing was conducted, covering cross-sectional dependence, panel serial correlation, groupwise heteroskedasticity, and cluster-robust standard errors. There are three approach models in panel data regression, namely the Common Effect Model (pooled least squares), Fixed Effect Model (One-Way & Two-Way), and Random Effect Model (RE). To account for macroeconomic dynamics over the 2018–2024 period including the COVID-19 pandemic, the revision of Law No. 16/2019 on marriage age limits, inflation,

and post-pandemic recovery, the Fixed Effects model is estimated under Unit Fixed Effects (One-Way) and Two-Way Fixed Effects (Cross-Section and Period/Time Effects) specifications to control for potential bias from time-varying common shocks.

$$\text{Divorce}_{it} = \beta_1 \text{EM}_{it} + \beta_2 \text{MYS}_{it} + \beta_3 \text{Pov}_{it} + \alpha_i + \lambda_t + \varepsilon_{it}$$

To determine which method to use among these three models, the Chow Test (Redundant Fixed Effects Tests) and Hausman Test were conducted, aiming to evaluate which model provides the best estimation for the research data. Furthermore, hypothesis testing was carried out, which includes the statistical t-test, statistical F-test, and the coefficient of determination R^2 .

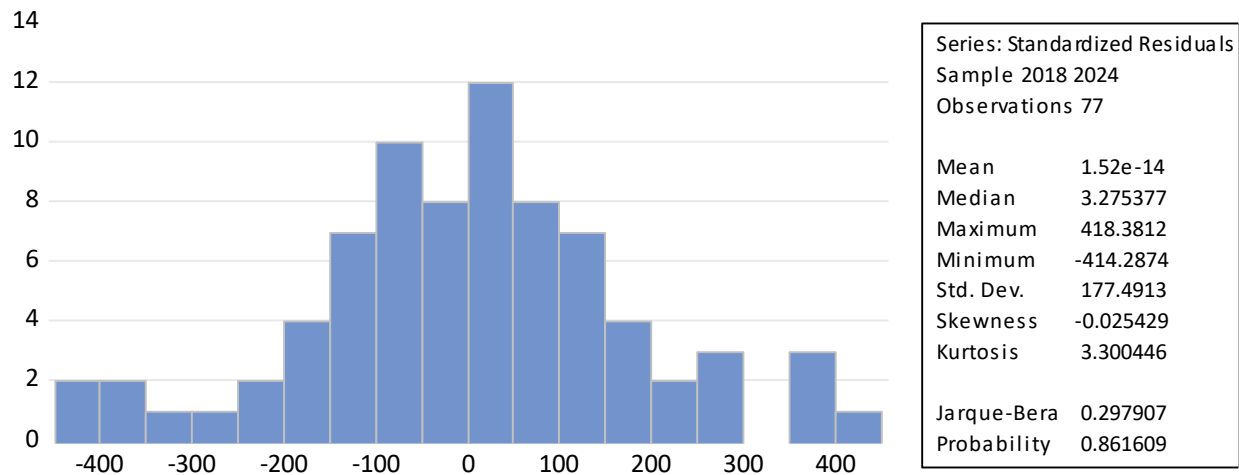
RESULT AND DISCUSSION

This research presents several statistical test results using E-Views. To ensure that the generated model can serve as a robust and efficient estimation tool, it must satisfy the classical assumption tests, which consist of the Normality Test, Multicollinearity Test, and Heteroscedasticity Test. Based on the data processing and analysis

conducted, the classical assumption test results are obtained as follows:

1. Normality Test

Figure 1: Normality Test Graph



Source: Data Processing Results for Normality Test

Based on Figure 1, it is known that the normality assumption of the residuals is satisfied through the Jarque-Bera statistical test. The results show a Jarque-Bera value of 0.297907 with a p-

value of 0.861609 (greater than 0.05), which means the data is normally distributed.

2. Multicollinearity Test

Table 2: Multicollinearity Test

	Pov	RLS	UPP
Pov	1.000000	0.390279	-0.242497
RLS	0.390279	1.000000	-0.136880
UPP	-0.242497	-0.136880	1.000000

Based on Table 2, it can be explained that all correlation coefficients of the independent variables are below 0.85. Therefore, it can be concluded that

the independent variables are free from multicollinearity issues.

Based on the correlation matrix between the research variables, the relationship can be detailed as follows:

- a. Poverty (Pov) and Mean Years of Schooling (MYS): Have a correlation coefficient of 0.390279. This value indicates a positive relationship with a low-to-moderate correlation strength. This means that an increase in poverty levels tends to be followed by an increase in mean years of schooling within the observed data, although the relationship is not very strong.
- b. Poverty (Pov) and Early marriage (EM): Have a correlation coefficient of -0.242497, indicating a negative relationship with a low correlation strength. This indicates that higher poverty levels correspond to a decreasing trend in the early marriage within the research data, though the relationship is weak, meaning changes in one variable do not substantially affect the other.
- c. Mean Years of Schooling (MYS) and Early marriage (EM): Have a correlation coefficient of -0.136880, which also indicates a negative relationship with a very low level of correlation. This implies that an increase in mean years of schooling

tends to be followed by a decrease in the early marriage, but the relationship is very weak, making its effect relatively minor.

3. Cross-Section Heteroskedasticity Test Test/ Groupwise Heteroskedasticity

Table 3: Heteroscedasticity Test

Heteroscedasticity Test	Glejser Test
Obs*R-squared	0.4284
Prob. Chi-Square(3)	0.9343*
Conclusion	Homoscedasticity

* = *significant, greater than alpha = 0.05*

Source: Data Processing Results for Heteroscedasticity Test

The Obs*R-squared value yields a Chi-Square probability of 0.9343 (> 0.05), demonstrating that the model exhibits homoskedasticity.

3. Autocorrelation Test and Serial Correlation Panel

Table 4: Autocorrelation Test

Mean dependent var	770.3377
S.D. dependent var	592.2593
Akaike info criterion	13.41304
Schwarz criterion	13.83918
Hannan-Quinn criter.	13.58349
Durbin-Watson stat	1.629834

Based on the autocorrelation test results in Table 4, observed from the Durbin-Watson (DW) statistic value,

a value of 1.629834 is obtained. Generally, the Durbin-Watson test criteria state that a DW value close to 2 indicates the absence of autocorrelation in the regression model. A DW value within the range of 1.5 to 2.5 is also frequently used as an indication that the model is free from autocorrelation.

As a robustness check and to satisfy advanced econometric diagnostic standards, the model was re-estimated using Cluster-Robust Standard Errors (White Cross-Section/Period Clustered SE) at the regency/city level to correct for potential heteroskedasticity and spatial/cross-sectional dependence.

4. Robustness Check: Estimation Using Cluster-Robust Standard Errors (White Cross-Section)

Table 5: Robustness Test Estimation Results (Cluster-Robust Standard Errors)

Variabel	Koefisien (Coefficient)	Std. Error	t-Statistic	Prob. (p-value)	Keterangan
Constanta (C)	1139.977	597.1135	1.909146	0.1048	Not Significant
Poverty	59.11882	32.01712	1.846475	0.1143	Not Significant
MYS	-129.0026	78.10248	-1.651709	0.1497	Not Significant
Early Marriage	1.643461	0.473056	3.474135	0.0132**	Significant ($\alpha = 5\%$)

Table 6: Characteristic & Statistic Model

Statistical Measure	Value	Statistical Measure	Value
R-squared	0.998604	Mean dependent var	668.9481
Adjusted R-squared	0.998139	S.D. dependent var	543.8175
S.E. of regression	23.45831	Akaike info criterion	9.367053
Sum squared resid	31366.65	Schwarz criterion	9.975833
Log likelihood	-340.6315	Hannan-Quinn criter.	9.610560
F-statistic	2146.676	Durbin-Watson stat	2.008148
Prob (F-statistic)	0.000000	Number of Observation	77 (11 CS $\times$ 7 Periode)

** indicates statistical significance at the 5% significance level ($p < 0.05$).

Effects: The model specification includes Cross-Section Fixed Effects and Period Fixed Effects (*Two-Way Fixed Effects*).

Consistency of Early Marriage Variable:

The Early Marriage variable was found

to be highly robust to both heteroskedasticity and spatial/cross-sectional correlation. The coefficient is positive at 1.6435 and remains statistically significant at the 5% significance level ($p = 0.0132 < 0.05$). This indicates that an increase in early marriage consistently contributes to an increase in the number of divorces.

The Poverty and MYS Variable: After adjusting the *standard errors* using clustering at the regency/city level, the Poverty variable ($p = 0.1143$) and MYS variable ($p = 0.1497$) were no longer statistically significant at the 5%

significance level, although Poverty remained marginally significant at the 10% level.

Goodness of Fit: The R^2 value of 0.9986 indicates that 99.86% of the variation in the number of divorces can be explained by the variables included in the model, along with regional and time effects (two-way fixed effects). The Durbin-Watson statistic of 2.0081 also suggests that the model is free from autocorrelation.

Model Selection Test

Table 7: Chow Test Result (Two-Way Redundant Fixed Effects)

Effects Test	Statistic	d.f.	Prob. (p-value)
Cross-section F	685.300961	(10, 57)	0.0000
Period F	0.444581	(6, 57)	0.8458
Cross-Section/Period Combined F	732.055846	(16, 57)	0.0000

Based on Table 7, the probability values for Cross-section F and Combined F are 0.0000 (< 0.05), leading to the rejection of

H_0 and demonstrating that the Fixed Effect Model (FEM) is superior to the CEM.

Table 8: Hausman Test Result

Test Summary	Chi-Sq. Statistic	Chi-Sq. d.f.	Prob. (p-value)
Cross-section random	31.960226	3	0.0000

Based on Table 8, the Chi-Square value for the Hausman test is 31.9602, with a p-value of 0.0000 (< 0.05). The null hypothesis (H0) is rejected and the

alternative hypothesis (H1) is accepted; thus, the Fixed Effect Model (FEM) is definitively selected as the best model.

Regression Analysis and Hypothesis Testing

Table 9: Estimation Results

Variable	Koefisien	t-Statistic	Prob.
C	5376,632	3,079611	0,0031
Pov	468,8395	2,965500	0,0043
EM	2,092257	3,991505	0,0002
MYS	-281,0761	-2,161505	0,0345

Based on the regression estimation results in Table 6, the regression equation model can be formulated as follows:

$$\text{PCR} = 5376.632 + 2.092257 \text{ EM}_{it} - 281.0761 \text{ MYS}_{it} + 468.8395 \text{ Pov}_{it}$$

1. Constant: Based on the regression results, the constant value is 5376.632. This indicates that the number of divorces cases unaffected by the early marriage, mean years of schooling, and poverty is 5376.632 cases. In other words, this serves as the baseline of divorce without any influence from the independent variables.
2. The Influence of Poverty on the Number of divorces: Based on the t-

statistic in Table 6, the poverty variable has a positive effect on divorce cases. When poverty increases by one percent, the number of divorcess will increase by 468.8395 cases, *ceteris paribus*. Theoretically, these findings are consistent with the financial stress mechanism. Poor households have more limited economic buffers when facing job loss, rising prices, or unexpected expenses. Pressure on household budgets may intensify conflicts over spending and the division of responsibilities. In accordance with Social Stress Theory and the Economic Hardship Model, economic constraints trigger

psychological strain and conflict regarding household financial provision, ultimately leading to divorce filings. However, this relationship may also be influenced by other factors, such as unemployment, employment structure, and household characteristics that are not included in the model.

3. The Influence of Early marriage on the number of divorces: The early marriage variable has a positive coefficient of 2.092257 with $p=0.0002$. This means that, in the estimated specification, one additional case of early marriage is associated with an increase of approximately 2.09 divorce cases, *ceteris paribus*. These findings are consistent with the mechanisms linking early marriage to household instability, including lower emotional readiness, incomplete education, limited life experience, economic pressures, and childcare responsibilities at a young age. The observed relationship between early marriage and marital dissolution is also supported by

existing evidence; however, the literature employing stronger identification strategies cautions that simple correlations may be influenced by unobserved factors. Therefore, the findings of this study are interpreted as a relationship at the regional level rather than as an estimate of individual-level causality.

4. The Influence of Mean Years of Schooling (MYS) on the number of divorces: The mean years of schooling has a coefficient of -281.0761 and is significant at the 5 percent level ($p=0.0345$). This means that a one-year increase in the mean years of schooling is associated with a decrease of approximately 281.08 divorce cases in the model, *ceteris paribus*. The negative direction can be explained by increased individual capacity to access information, plan families, communicate effectively, and manage conflicts, as well as by greater economic opportunities. Drawing on Human Capital Theory and New Home Economics (Becker, 1981), education enhances cognitive

maturity, emotional conflict management, and family planning stability, thereby reducing the incidence of divorce. However, the relationship between education and divorce is not necessarily negative. Higher education may also increase economic independence, expand alternatives outside marriage, and reduce the economic costs associated with ending an unstable marriage. Therefore, the negative finding in the West Kalimantan study should be understood as an empirical finding specific to the particular context and period, rather than as a general rule that education always reduces divorce.

Goodness of Fit Tests

Table 10 : Simultaneous significance test

Indikator	Nilai
F-statistic	56.80411
Prob(F-statistic)	0.000000
R-squared	0.921393
Adjusted R-squared	0.905172
R-Squared	0.921393
Adjusted R-Squared	0.905172

Based on Table 10, The F-statistic of 56.80411, with a Prob (F-statistic) of

0.000000, indicates that poverty, early marriage, and mean years of schooling are jointly significant in explaining the number of divorcess in the estimated fixed-effects specification. The R-squared value of 0.921393 indicates that the overall goodness-of-fit of the model is 92.14 percent. The Adjusted R-squared value of 0.905172 indicates that the model maintains a high goodness-of-fit after adjusting for the number of parameters. Based on the results of the analysis examining the effect of poverty on the number of divorces in West Kalimantan Province, poverty was found to have a positive and statistically significant effect on the number of divorces, with a t-statistic of 2.965500 and a probability value of 0.0043. These findings indicate that, statistically, poverty is significantly associated with the number of divorces in West Kalimantan Province. The positive coefficient suggests that an increase in the poverty rate tends to be accompanied by an increase in the number of divorces. Based on the panel data regression results, the early marriage variable was found to have a positive and statistically

significant effect on the number of divorces in West Kalimantan Province. This is indicated by a t-statistic of 3.991505 and a probability value of 0.0002, which is lower than the 0.05 significance level. These results indicate that the hypothesis stating that early marriage has a positive effect on the divorce is supported. The positive coefficient indicates that a higher prevalence of early marriage tends to be associated with a higher likelihood of divorce. The results of the study show that mean years of schooling (MYS) has a negative and statistically significant effect on the divorce in West Kalimantan Province, with a t-statistic of -2.161505 and a probability value of 0.0345, which is below the 0.05 significance level. These findings indicate that the hypothesis stating that mean years of schooling has an effect on the divorce is supported. The negative coefficient indicates that higher mean years of schooling among the population is associated with a lower tendency toward divorce in West Kalimantan Province.

CONCLUSION

Based on the results of data analysis and discussion, this study provides conclusions regarding the relationship between poverty, early marriage, mean years of schooling, and divorce in West Kalimantan Province as follows:

1. The poverty variable has a positive and significant effect on the divorce in West Kalimantan Province. Unstable economic conditions can create pressure within the household, thereby triggering conflicts between husbands and wives that can ultimately lead to divorce.
2. The variable of early marriage has a positive and significant effect on the number of divorces in West Kalimantan Province. This result indicates that the higher the number of marriages at a young age, the higher the tendency for divorces to occur. Marriages at an early age are often confronted with various challenges, such as emotional unpreparedness, limited life experience, and unstable economic conditions.

3. The mean years of schooling variable has a negative and significant effect on the number of divorces in West Kalimantan Province. This result indicates that a higher level of public education, as reflected by the mean years of schooling, tends to decrease the number of divorces. Education can enhance an individual's capacity to understand roles and responsibilities within the household, as well as assist couples in resolving conflicts more rationally.

Simultaneously, the variables of poverty, age at first marriage, and mean years of schooling have a significant effect on the number of divorces in West Kalimantan Province. This demonstrates that poverty, early marriage, and mean years of schooling serve as good regressors for divorce cases. Furthermore, the R-squared value of 0.921393 indicates a high overall goodness-of-fit of the fixed-effects model. However, this value should not be interpreted as indicating that poverty, early marriage, and mean

years of schooling alone explain 92.14% of the variation in the number of divorces, as the R-squared also captures the contribution of the fixed effects across regencies/cities.

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